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Probiotics for the prevention of gestational diabetes mellitus: A meta-analysis of randomized controlled trials

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DOI: https://doi.org/10.17305/bb.2024.10339

Conflicts of interest: Authors declare no conflicts of interest.

Submitted: 16 February 2024 / Accepted: 6 April 2024 / Published: 19 April 2024

Funding: No funding was received for this study.

Data availability: All data generated during the study are presented in the manuscript.
ABSTRACT

Changes of intestinal microbiota have been shown to be involved in the development of gestational diabetes mellitus (GDM). We performed a meta-analysis to systematically evaluate the potential role of probiotics for the prevention of GDM. Systematic literature search was performed in electronic databases including PubMed, Cochrane library, Embase, Web of Science, Wanfang, and CNKI to obtain relevant randomized controlled studies. A random-effects model was used to pool the results by incorporating the impact of the potential heterogeneity. Meta-regression and subgroup analyses were conducted to evaluate the source of heterogeneity. Fourteen studies involving 3527 pregnant women were included. Results showed that probiotics significantly reduced the incidence of GDM as compared to control (risk ratio [RR]: 0.71, 95% confidence interval: 0.52–0.96, \( P = 0.03 \)) with significant heterogeneity (\( I^2 = 73\% \)). The meta-regression showed that body mass index (BMI) of females was positively associated with the RR for the effect of probiotics on GDM (coefficient = 0.084, \( P = 0.01 \)). The results of subgroup analyses also suggested that probiotics significantly reduced the risk of GDM in women with BMI < 26 kg/m\(^2\), but not in those with BMI \( \geq 26 \) kg/m\(^2\) (\( P \) for subgroup difference = 0.001). In addition, the preventative efficacy of probiotics on GDM was remarkable in women < 30 years, but not in those \( \geq 30 \) years (\( P \) for subgroup difference < 0.001). In conclusion, probiotics may be effective in reducing the risk of GDM, particularly for females with lower BMI and younger age.

Keywords: Gestational diabetes mellitus (GDM); probiotics; prevention; incidence; meta-analysis
INTRODUCTION

Gestational diabetes mellitus (GDM) is a prevalent metabolic disorder that occurs during pregnancy (1, 2). Existing literature suggests that the prevalence of GDM among pregnant individuals ranges from 15% to 20% (1). Risk factors associated with GDM include advanced maternal age, elevated body mass index (BMI), familial history of type 2 diabetes mellitus (T2DM), and a prior history of GDM in previous pregnancy (3). Emerging research indicates that GDM is not only linked to immediate adverse outcomes such as miscarriage, preterm birth, and macrosomia (4, 5), but it is also associated with a range of long-term health risks for both mothers and their offspring, including maternal and child obesity, increased risk of type 2 diabetes, and heightened maternal susceptibility to cancer and cardiovascular diseases (4, 6, 7). Consequently, there is a pressing need for the development of innovative approaches to prevent the onset of GDM (8).

Pregnancy has been associated with disruptions in the homeostasis of intestinal microbiota, with a notable increase in actinobacteria and proteobacteria observed in 60 to 70% of women (9, 10). Studies have shown that women with GDM exhibit more pronounced alterations in gut microbiota compared to those without GDM, resembling patterns seen in non-pregnant women with T2DM (11, 12). This suggests a potential role of gut microbiota in the development of GDM. Probiotics, as living microorganisms, play a beneficial role in restoring and maintaining the balance of gut microbiota composition (13). In T2DM patients, the use of probiotics has been linked to a reduction in insulin resistance and enhancement of glycemic control (14, 15). Furthermore, in females with a confirmed diagnosis of GDM, supplementation with probiotics has demonstrated improvements in hyperglycemia and dyslipidemia, as well as a decrease in birth weight of their offspring (16-18). Similarly, probiotics supplementation has been suggested to improve glycemic control via multiple mechanisms, such as reducing inflammation, enhancing the production of short-chain fatty acids (SCFAs), regulation of gut microbiota, improving insulin
sensitivity, and preventing excessive weight gain (19, 20). However, conflicting findings arise from previous studies examining the efficacy of probiotics in preventing GDM (21). Two meta-analyses conducted previously did not find significant evidence to support the use of probiotics in reducing the risk of GDM (22, 23). However, they included only five to six studies and significant heterogeneity, which were not explored due to the limited number of available studies, was observed in both (22, 23). Additional randomized controlled trials have been published since (24-29). Accordingly, the aim of our study was to perform an updated meta-analysis to comprehensively evaluate the influence of probiotics supplementation of the incidence of GDM in pregnant women.
MATERIALS AND METHODS

This study is in accordance with the guidelines of Preferred reporting items for systematic reviews and meta-analyses (PRISMA) (30, 31) and Cochrane Handbook (32).

Study inclusion and exclusion criteria

The principle of PICOS, which is explained below, was utilized to determine the inclusion criteria for the meta-analysis.

P (participants): Women planning to conceive or at early pregnancy; I (intervention): Probiotics supplements during pregnancy, with no restrictions to the strains, timing, or dose of probiotics; C (control): Placebo or no additional treatment; O (outcomes): Reported the incidence of GDM during follow-up. The methods and criteria for the diagnosis of GDM were in accordance with that reported in the original studies. S (study design): Only RCTs with parallel groups that were published as complete articles in English or Chinese in peer-reviewed journals were deemed eligible for study design. Non-randomized studies, studies not including women planning to conceive or at early pregnancy, not with an intervention of probiotic supplementation, or not reporting the outcome of GDM incidence were excluded. In case studies with potentially overlapping patient populations were found, the meta-analysis included the one that had the larger sample size.

Literature search strategy

To identify studies in Medline (PubMed), CENTER (Cochrane Library), Embase (Ovid), Web of Science, Wanfang, and China National Knowledge Infrastructure (CNKI), a search strategy was employed that encompassed the following criteria by combination of the keywords: (1) "probiotic" OR "probiotics" OR "lactobacillus" OR "lactobacilli" OR "bifidobacteria" OR "bifidobacterium"; (2) "gestational diabetes mellitus" OR "GDM" OR (("gestational" OR "pregnancy" OR "pregnant") AND ("diabetes" OR "diabetic" OR "hyperglycemia")); and (3) "random" OR "randomized" OR "randomised" OR "randomly"
OR "allocated" OR "control" OR "placebo". Our focus was solely on research that involved human participants. In addition, we conducted a manual search for references to relevant reviews and primary articles. The most recent database search was conducted on December 21, 2023.

**Extraction of data and assessment of study quality**

Two authors conducted separate searches in databases, gathered information, and assessed the quality. In case of any disagreements, the corresponding author was consulted to reach a consensus. For the study, various data were gathered including general details, characteristics of the study design, participants characteristics, age, body mass index (BMI), proportions of women with primipara, use of lifestyle recommendations (diet and exercise), details of interventions (probiotics used, timing, and dose), regimens of controls, and criteria for the diagnosis of GDM. Cochrane’s Risk of Bias Tool (32) was used to evaluate the quality of RCTs included in this review. It assessed seven domains, including the generation of random sequence, concealment of allocations, blinding of participants and personnel, blinding of outcome evaluation, incomplete result data, and selective reporting of outcomes.

**Statistical analysis**

The incidence of GDM, compared between women with probiotics supplementation and women in the control group, was summarized as risk ratio (RR) and corresponding 95% confidence interval (CI). The outcome data was extracted using the intention-to-treat principle. The Cochrane Q test was used to investigate the heterogeneity among included studies (32). Furthermore, the $I^2$ statistic was calculated, where $I^2 > 50\%$ suggested statistical heterogeneity (33). To incorporate potential heterogeneity, a random-effects model was employed for pooling the data (32). For outcomes of adequate datasets (10 or above), meta-regression and subgroup analyses according to study characteristics were
performed to evaluate the source of heterogeneity. The meta-regression analysis tested the significance of the individual study characteristics’ influence on the results of the meta-analysis, with a $P$ value $< 0.05$ indicating a significant modification effect. A positive coefficient demonstrated that the evaluated study characteristics are positively related to the OR of the results, while a negative coefficient demonstrated that the evaluated study characteristics are negatively related to the OR of the results. These characteristics included study country, mean age, BMI, timing, and dose of probiotics supplementation, and the risk of GDM of the studied females as reflected by the incidence of GDM in the control groups. Medians of continuous variables were selected as the cutoffs to define the subgroups. Publication bias was evaluated using Egger's test for regression asymmetry and funnel plots (34). A $P$-value $< 0.05$ suggested a statistically significant distinction. The statistical analysis was conducted using RevMan (Version 5.1; Cochrane, Oxford, UK) and Stata (Version 12.0; Stata Corporation, US) software.
RESULTS

Literature search

The process of acquiring literature is illustrated in Figure 1. In summary, a total of 719 articles were obtained through database searches, with 530 remaining after removing duplicates. A total of 494 articles were subsequently excluded by screening via titles and abstracts, primarily because they were not relevant to the objective of the study. After reading the full text, an additional 22 articles out of the initial 36 were excluded due to the reasons outlined in Figure 1. At last, fourteen RCTs (24-29, 35-42) were available for the subsequent meta-analysis.

Study characteristics and data quality evaluation

Table 1 provides a summary of the studies included in the meta-analysis. In total, there were 14 RCTs involving 3527 females who were planning to conceive in upcoming six months or at early pregnancy (24-29, 35-42). These studies were published between 2010 and 2022, and carried out in Finland, Ireland, New Zealand, Australia, China, Denmark, Iran, the United Kingdom, Singapore, and Pakistan. The mean ages of the females were 27 to 34 years, and the mean BMI scores were 21 to 39 kg/m². The proportions of females with primipara varied between 15.0 to 63.5%. In four studies, dietary recommendation was also provided to females of the intervention and control groups (25, 35, 36, 41). However, no evaluation has been performed regarding the diet or physical activities pre and post-intervention among these studies. Multiple different strains were used for probiotics supplementation, such as Lactobacillus rhamnosus, Lactobacillus acidophilus, Lactobacillus salivarius, Bifidobacterium lactis, Bifidobacterium longum, Bifidobacterium bifdum, Lactobacillus casei, Lactobacillus bulgaricus, Lactobacillus plantarum, Lactobacillus paracasei, Bifidobacterium breve, Bifidobacterium infantis, and Streptococcus thermophilus, with Lactobacillus rhamnosus GG as the most commonly used
probiotics strain. Most of the included studies used multiple strains as intervention except for three studies (36-38), in which single-strain probiotics were used. The timing for the starting of probiotics supplementation varied among the included studies, ranging from within the first trimester to the gestational age (GA) of 24 weeks. The total doses of probiotics were 1 to 50×10⁹ colony-forming units per day. As for the controls, placebo capsules were used in 12 studies (26-29, 35-42), while for the other two studies, no additional treatment was considered as controls (24, 25). The incidence of GDM was diagnosed with the International Association of Diabetes in Pregnancy Study Group criteria (43) in all the studies using a “one-step” 2-hour 75 g oral glucose tolerance test (OGTT) except for one study (36), in which GDM was diagnosed with the American College of Obstetricians and Gynecologists criteria using a “two-step” 3-hour 100 g OGTT test (44). Compliance data were reported in three studies, with similar mean adherence rates of 94.5% (37), 88.4% (40), and > 90% (38) between females of the intervention and control groups, indicating good compliance. The incidence of adverse events was reported in two studies (26, 40). Only mild discomfort related to the treatments was reported, which was similar in females in the intervention and control groups, with gastrointestinal symptoms being the most common symptoms.

Table 2 provides a detailed analysis of the included RCTs using Cochrane’s Risk of Bias Tool. One of the included study was open-label (25), another one was single-blinded (24), while the remaining 12 studies were double-blinded (26-29, 35-42). The details of the random sequence generation were reported in nine studies (26, 27, 36-42), and seven studies reported the details of allocation concealment (27, 36, 38-42).

**Influence of probiotics on the incidence of GDM**

Pooled results of 14 studies using random-effects models showed that probiotics significantly reduced the incidence of GDM as compared to control (RR: 0.71, 95% CI: 0.60-0.84).
0.52 – 0.96, \( P = 0.03 \); Figure 2A) with significant heterogeneity (\( I^2 = 73\% \)). Sensitivity analysis by excluding the study with ACOG criteria (36) for the diagnosis of GDM retrieved similar results (RR: 0.70, 95% CI: 0.51 – 0.95, \( P = 0.02 \); \( I^2 = 75\% \)). In addition, sensitivity analysis excluding the two studies with probiotics started at the 24 weeks of GA (36, 41) also showed similar results (RR: 0.71, 95% CI: 0.51 – 0.98, \( p = 0.04 \); \( I^2 = 76\% \)).

The meta-regression showed that the females’ mean BMI was positively associated with the RR for the effect of probiotics on GDM (coefficient = 0.084, \( P = 0.01 \); Figure 2B and Table 3), which largely explained the source of between-study heterogeneity (residual \( I^2 = 10.5\% \)). Other variables such as sample size, mean age, probiotics dose, median GA for starting probiotics, or incidence of GDM in control groups were not suggested to be significant modifiers for the effect of probiotics on GDM, according to the results of the meta-regression analyses (\( P \) all > 0.05, Table 3).

Subsequent subgroup analyses according to study country did not significantly affect the results (\( P \) for subgroup difference = 0.09; Figure 3A). However, the results of subgroup analyses indicated that the preventative efficacy of probiotics on GDM was remarkable in females < 30 years, but not in those \( \geq 30 \) years (RR: 0.42 versus 1.05, \( P \) for subgroup difference < 0.001; Figure 3B). In addition, it was also indicated that probiotics significantly reduced the risk of GDM in females with BMI < 26 kg/m\(^2\), but not in those with BMI \( \geq 26 \) kg/m\(^2\) (\( P \) for subgroup difference = 0.001; Figure 4A). Subgroup analyses did not support that other study characteristics could significantly influence the effect of probiotics supplementation on the risk of GDM, such as probiotics dose (\( P \) for subgroup difference = 0.70; Figure 4B), timing of probiotic supplementation (\( P \) for subgroup difference = 0.53; Figure 5A), or risk of GDM as reflected by the incidence of GDM in controls (\( P \) for subgroup difference = 0.97; Figure 5B).

**Publication bias**
The funnel plots for the meta-analysis of the probiotics’ influence on the incidence of GDM in pregnant women are shown in Figure 6. The funnel plots are symmetrical on visual inspection, suggesting the low risk of publication bias. The results of Egger’s regression test also suggested a low risk of publication bias ($P = 0.39$).
DISCUSSION

In our study, by pooling the results of 14 RCTs, we found that probiotics supplementation during pregnancy could significantly reduce the incidence of GDM. Interestingly, subsequent meta-regression and subgroup analyses suggested that the BMI of the pregnant females may significantly modify the effect of probiotics on GDM, which largely explained the source of heterogeneity. Specifically, probiotics significantly reduced the risk of GDM in women with BMI < 26 kg/m\(^2\), but not in those with BMI ≥ 26 kg/m\(^2\). In addition, the preventative efficacy of probiotics on GDM was remarkable in women < 30 years, but not in those ≥ 30 years. Taken together, results of this meta-analysis indicate that probiotics may be effective in reducing the risk of GDM, particularly for females with lower BMI and younger age.

Although several meta-analyses have been published on the topic of influence of probiotics supplementation on the risk of GDM (22, 23), this current updated meta-analysis has several methodological strengths compared to the previous ones. First, in this meta-analysis, we performed extensive literature search in six commonly used electronic databases, and retrieved eligible RCTs which investigated the efficacy of probiotics for the prevention of GDM. As a result, fourteen studies involving 3527 pregnant females were included. The overall sample size of the meta-analysis is much larger than that of the previous ones (22, 23). In addition, multiple meta-regression and subgroup analyses were performed to identify the study characteristics’ influences on the outcome and to determine the source of heterogeneity.

We found that the BMI of the females was positively associated with the RR of the probiotics effect on GDM, and probiotics significantly reduced the risk of GDM in females with BMI < 26 kg/m\(^2\), but not in those with BMI ≥ 26 kg/m\(^2\). Similarly, the preventative efficacy of probiotics on GDM was remarkable in women < 30 years, but not in those ≥ 30
years. The mechanisms underlying these findings remain to be determined. Interestingly, it has been confirmed that advanced maternal age (45) and obesity (46) are established risk factors of GDM. Therefore, findings of this study may suggest that probiotics supplementation is effective to reduce the risk of GDM in low-risk women, but not in high-risk women. Physiologically, the mechanisms underlying the effects of probiotics supplementation during pregnancy is to attenuate the gut dysbiosis related to pregnancy, a potential pathway linked to the pathogenesis of GDM (47). For females with high risk for GDM, multiple mechanisms may be involved in the pathogenesis of GDM besides dysregulation of intestinal microbiota, such as islet beta-cell dysfunction, insulin resistance, neurohormonal dysregulation, oxidative stress, and inflammation (48), and probiotics supplementation may become less effective. In addition, a previous study suggested that multi-strain probiotics are beneficial for improved metabolic and inflammatory outcomes in post-GDM women by modulating gut dysbiosis, which highlighted the necessity for a comprehensive strategy for postpartum treatment that includes probiotics to protect post-GDM women from developing glucose intolerance (49). Accordingly, females with advanced age and obesity may respond poorly to probiotics because they have lesser gut microbial diversity. Moreover, a recent meta-analysis indicated that probiotics may have positive effects on metabolic, inflammation, oxidative stress, and neonatal outcomes in females with GDM. Additionally, diet and pre-intervention washout may modify the effects of probiotics (50). These factors may also confound the efficacy of probiotics in pregnant females with high risk for GDM. These hypotheses should be validated in future studies.

This meta-analysis has limitations. First, the species/strains of probiotics varied among the included studies, which may also lead to heterogeneity. Future studies should be performed to determine the optimal species/strains for the prevention of GDM. Second, although our meta-regression and subgroup analyses did not show that differences in dose or timing for starting probiotics supplementation may modify the effect of probiotics for
prevention of GDM, the optimal dose and timing for starting probiotics remain to be clarified in this clinical setting. Third, the incidence of GDM could be significantly affected by dietary habits and physical activates (51), two key factors which may modify the potential preventative efficacy of probiotics on GDM. However, these two factors were rarely reported or controlled among the included studies. In addition, most of the studies did not evaluate the baseline gut microbial diversity and did not observe the effect of probiotics on gut microbial diversity after intervention. Finally, this meta-analysis was based on study-level data rather than individual patient-level data. Accordingly, results of the meta-regression and subgroup analyses should be interpreted with caution. Large-scale RCTs are still needed to validate these findings.

CONCLUSION

Taken together, probiotics supplementation may be effective in reducing the risk of GDM, particularly for females with lower BMI and young age. Although the optimal species/strains, dose, and starting timing of probiotics supplementation remain to be determined, these findings support the potential use of probiotics supplementation as an effective strategy to reduce the incidence of GDM in pregnant females. Further research is needed to evaluate the influence of probiotic supplementation on the risk of GDM in high-risk females, such as those with advanced age and obesity, especially high-quality RCTs.
REFERENCES


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TABLES AND FIGURES WITH LEGENDS
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<tr>
<th>Study</th>
<th>Location</th>
<th>Design</th>
<th>Participants</th>
<th>Patient number</th>
<th>Mean age (years)</th>
<th>Mean BMI (kg/m²)</th>
<th>Primipara (%)</th>
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<th>Timing of intervention</th>
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<th>Total dose (10⁹ cfu/d)</th>
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<td>14 weeks</td>
<td>32</td>
<td>30.2</td>
<td>NR</td>
<td>No dietary counseling</td>
<td>Lactobacillus acidophilus LA1, Bifidobacterium longum sp54 cs, and Bifidobacterium bifidum sp9 cs</td>
<td>15</td>
<td>Placebo capsule</td>
<td>IADPSG criteria</td>
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<td>Godfrey 2021</td>
<td>United Kingdom, Singapore, and New Zealand</td>
<td>R, DB, PC</td>
<td>Women planning to conceive in upcoming 6 months</td>
<td>Before pregnancy to delivery</td>
<td>30.3</td>
<td>25.7</td>
<td>63.5</td>
<td>NR</td>
<td>Lactobacillus rhamnosus NCC 4007 and Bifidobacterium animalis subspecies lactis NCC 2818</td>
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<td>Placebo capsule</td>
<td>IADPSG criteria</td>
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<tr>
<td>Study</td>
<td>Country</td>
<td>Randomization</td>
<td>Type of Intervention</td>
<td>Number</td>
<td>Mean BMI</td>
<td>Median or Mean</td>
<td>Risk of GDM</td>
<td>Duration</td>
<td>Microbiota</td>
<td>Duration</td>
<td>Treatment</td>
<td>Criteria</td>
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<td>Baloch 2022</td>
<td>Pakistan</td>
<td>R, DB, PC</td>
<td>Women at early pregnancy with high risk for GDM</td>
<td>160</td>
<td>30</td>
<td>26</td>
<td>NR</td>
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<td>Second trimester (GA: 13~14 weeks) to delivery</td>
<td>Streptococcus, lactobacilli, and Bifidum</td>
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<td>Placebo capsule</td>
<td>IADPSG criteria</td>
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<tr>
<td>Liu 2022</td>
<td>China</td>
<td>R, DB, PC</td>
<td>Women at early pregnancy</td>
<td>112</td>
<td>29.7</td>
<td>22.6</td>
<td>NR</td>
<td>NR</td>
<td>Second trimester (GA: 20 weeks) to delivery</td>
<td>Streptococcus, lactobacilli, and Bifidum</td>
<td>12</td>
<td>Placebo capsule</td>
<td>IADPSG criteria</td>
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</tbody>
</table>

BMI, body mass index; cfu, colony-forming unit; GDM, gestational diabetes mellitus; R, randomized; DB, double-blinded; PC, placebo-controlled; OL, open-label; SB, single-blinded; GA, gestational age; IADPSG, the International Association of Diabetes in Pregnancy Study Group; ACOG, the American College of Obstetricians and Gynecologists; NR, not reported;
Table 2. Study quality evaluation via the Cochrane’s Risk of Bias Tool.

<table>
<thead>
<tr>
<th>Study</th>
<th>Random sequence generation</th>
<th>Allocation concealment</th>
<th>Blinding of participants</th>
<th>Blinding of outcome assessment</th>
<th>Incomplete outcome data addressed</th>
<th>Selective reporting</th>
<th>Other sources of bias</th>
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<td>Luoto 2010</td>
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<td>Shahriari 2021</td>
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Table 3. Results of univariate meta-regression analysis.

<table>
<thead>
<tr>
<th>Variables</th>
<th>RR for the incidence of GDM</th>
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</thead>
<tbody>
<tr>
<td></td>
<td>Coefficient</td>
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<tr>
<td>Sample size</td>
<td>0.0010</td>
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<tr>
<td>Mean age (years)</td>
<td>0.082</td>
</tr>
<tr>
<td>BMI (kg/m²)</td>
<td>0.084</td>
</tr>
<tr>
<td>Dose of probiotics (10⁹ cfu/d)</td>
<td>-0.00057</td>
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<td>Median GA for starting probiotics</td>
<td>0.0048</td>
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<tr>
<td>Incidence of GDM in control group (%)</td>
<td>-0.010</td>
</tr>
</tbody>
</table>

RR, risk ratio; GDM, gestational diabetes mellitus; CI, confidence interval; BMI, body mass index; cfu, colony-forming unit; GA, gestational age
Figure 1. Flowchart of the literature search.
Figure 2. Meta-analysis for the role of probiotics on the incidence of GDM in pregnant women. (A) Forest plots for the overall meta-analysis of the influence of probiotics on the incidence of GDM; and (B) univariate regression analysis for the influence of BMI on the efficacy of probiotics for the prevention of GDM.
Figure 3. Forest plots for the subgroup analyses of the efficacy of probiotics for the prevention of GDM. (A) Subgroup analysis based on country and (B) based on females’ mean ages.
Figure 4. Forest plots for the subgroup analyses of the efficacy of probiotics for the prevention of GDM. (A) Subgroup analysis based on the mean BMI of females and (B) subgroup analysis based on the probiotics dose.
Figure 5. Forest plots for the subgroup analyses of the efficacy of probiotics for the prevention of GDM. (A) Subgroup analysis based on the timing of probiotics supplementation and (B) based on the risk of GDM of the included females as reflected by the incidence of GDM in control groups.
Figure 6. Funnel plots evaluating the publication bias of the meta-analysis for the role of probiotics on the incidence of GDM in pregnant females.